### Next-generation statistical inference tools

Simulation-based inference, marginal statistics & accelerated nested sampling

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#### Simulation-based inference

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#### Accelerated nested sampling

Why can nested sampling be slow?

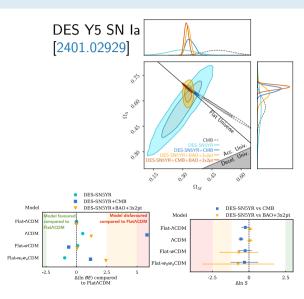
Accelerating with  $\beta$ -flows

Accelerating with jax

The standard approach if you are fortunate enough to have a likelihood function  $P(D|\theta)$ :

$$P(\theta|D) = \frac{P(D|\theta)P(\theta)}{P(D)}$$

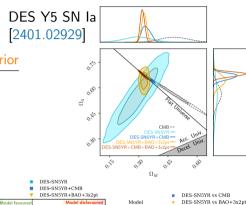
- 1. Define prior  $\pi(\theta)$ 
  - spend some time being philosophical
- 2. Sample posterior  $\mathcal{P}(\theta|D)$ 
  - use out-of-the-box MCMC tools such as emcee or MultiNest
  - make some triangle plots
- 3. Optionally compute evidence  $\mathcal{Z}(D)$ 
  - e.g. nested sampling or parallel tempering
  - do some model comparison (i.e. science)
  - talk about tensions

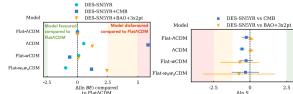


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 Posterior =  $\frac{\text{Likelihood} \times \text{Prior}}{\text{Evidence}}$ 

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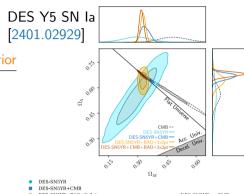


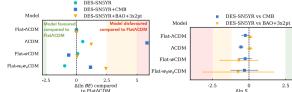


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$$\mathcal{P}(\theta|D) = \frac{\mathcal{L}(D|\theta)\pi(\theta)}{\mathcal{Z}(D)}$$
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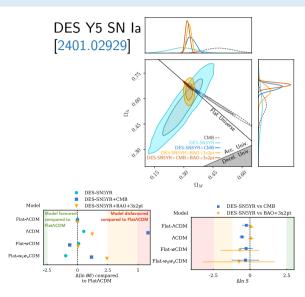




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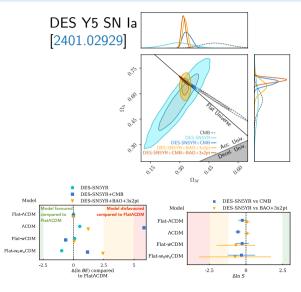
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$$\mathcal{P} \times \mathcal{Z} = \mathcal{J} = \mathcal{L} \times \pi$$
, Joint =  $\mathcal{J} = P(\theta, D)$ 

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# The three pillars of Bayesian inference

#### Parameter estimation

What do the data tell us about the parameters of a model? e.g. the size or age of a  $\Lambda CDM$ universe

$$P(\theta|D, M) = \frac{P(D|\theta, M)P(\theta|M)}{P(D|M)} \qquad P(M|D) = \frac{P(D|M)P(M)}{P(D)}$$

$$P = \frac{\mathcal{L} \times \pi}{2} \qquad \frac{\mathcal{Z}_{M}\Pi_{M}}{\sum_{m} Z_{m}\Pi_{m}}$$

$$Posterior = \frac{Likelihood \times Prior}{Likelihood \times Prior}$$

### Model comparison

How much does the data support a particular model? e.g.  $\Lambda CDM$  vs a dynamic dark energy cosmology

$$P(M|D) = \frac{P(D|M)P(M)}{P(D)}$$

$$\frac{\mathcal{Z}_{M}\Pi_{M}}{\sum_{m} Z_{m}\Pi_{m}}$$

 $Posterior = \frac{Evidence \times Prior}{Normalisation}$ 

### **Tension quantification**

Do different datasets make consistent predictions from the same model? e.g. CMB vs Type IA supernovae data

$$\mathcal{R} = \frac{\mathcal{Z}_{AB}}{\mathcal{Z}_{A}\mathcal{Z}_{B}}$$

$$\log \mathcal{S} = \langle \log \mathcal{L}_{AB} \rangle_{\mathcal{P}_{AB}} - \langle \log \mathcal{L}_{A} \rangle_{\mathcal{P}_{A}} - \langle \log \mathcal{L}_{B} \rangle_{\mathcal{P}_{B}}$$

# The three pillars of Bayesian inference

#### Parameter estimation

What do the data tell us about the parameters of a model? e.g. the masses and spins of a BBH collision

$$P(\theta|D,M) = \frac{P(D|\theta,M)P(\theta|M)}{P(D|M)} \quad P(M|D) = \frac{P(D|M)P(M)}{P(D)}$$

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### Model comparison

How much does the data support a particular model? e.g. IMRPhenom vs **FOBNR** waveform models

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### **Tension quantification**

Do different datasets make consistent predictions from the same model? e.g. Automated glitch detection

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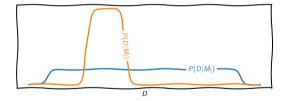
# Model comparison $\mathcal{Z} = P(D|M)$

Bayesian model comparison allows mathematical derivation of key philosophical principles.

Viewed from data-space *D*:

### Popper's falsificationism

- Prefer models that make bold predictions.
- ▶ if proven true, model more likely correct.



▶ Falsificationism comes from normalisation

Viewed from parameter-space  $\theta$ :

#### Occam's razor

- Models should be as simple as possible
- ...but no simpler
- Occam's razor equation:

$$\log \mathcal{Z} = \langle \log \mathcal{L} \rangle_{\!\mathcal{P}} - \mathcal{D}_{\mathsf{KL}}$$

• "Occam penalty": KL divergence between prior  $\pi$  and posterior  $\mathcal{P}$ .

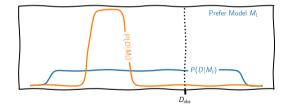
$$\mathcal{D}_{\mathsf{KL}} \sim \log \frac{\mathsf{Prior} \ \mathsf{volume}}{\mathsf{Posterior} \ \mathsf{volume}}$$

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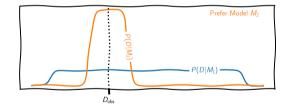
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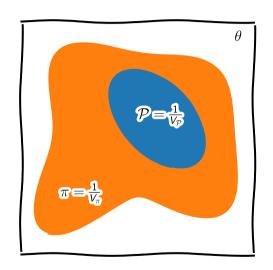
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# Why do sampling?

- ► The cornerstone of numerical Bayesian inference is working with **samples**.
- Generate a set of representative parameters drawn in proportion to the posterior  $\theta \sim \mathcal{P}$ .
- The magic of marginalisation ⇒ perform usual analysis on each sample in turn.
- The golden rule is stay in samples until the last moment before computing summary statistics/triangle plots because

$$f(\langle X \rangle) \neq \langle f(X) \rangle$$

• Generally need  $\sim \mathcal{O}(12)$  independent samples to compute a value and error bar.

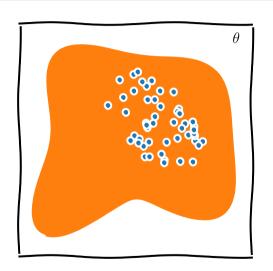


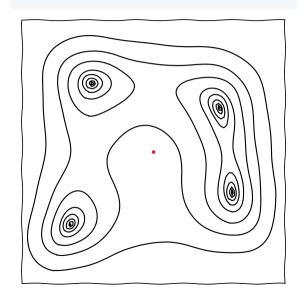
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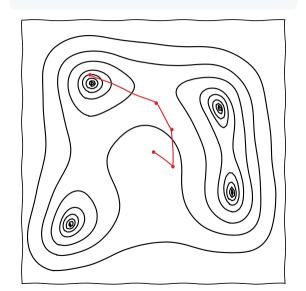
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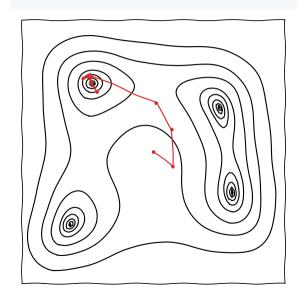
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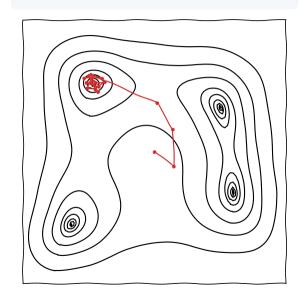
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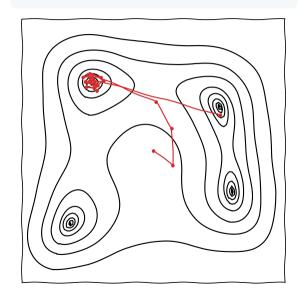


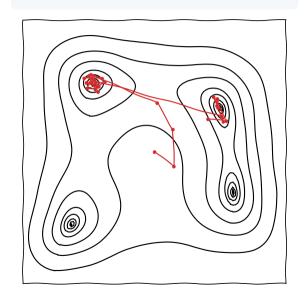


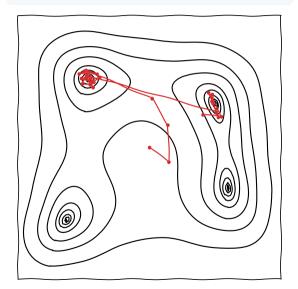


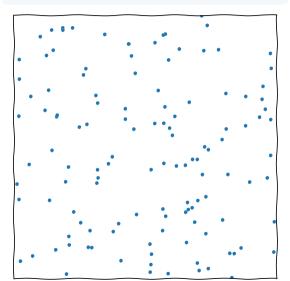


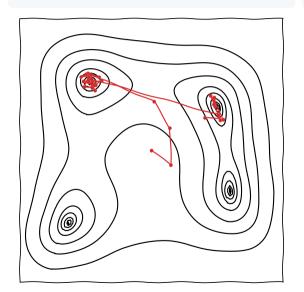


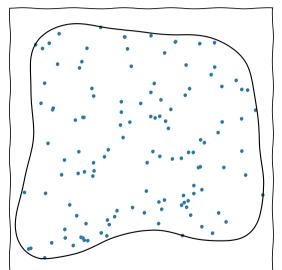


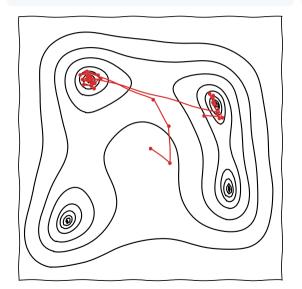


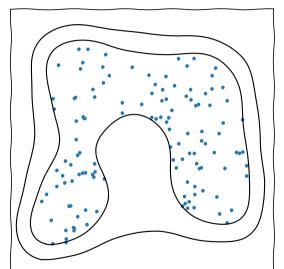


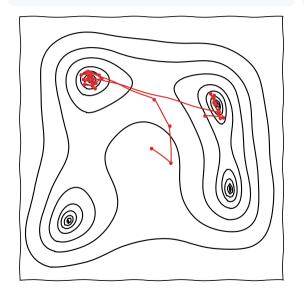


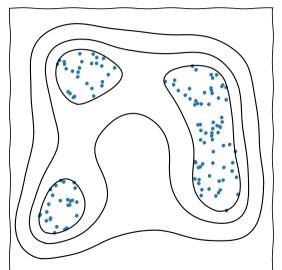


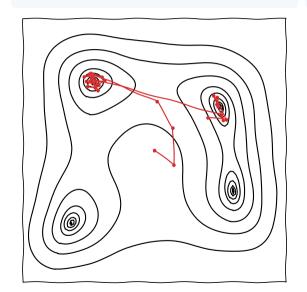


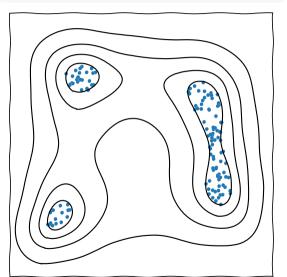


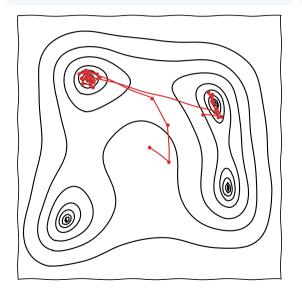


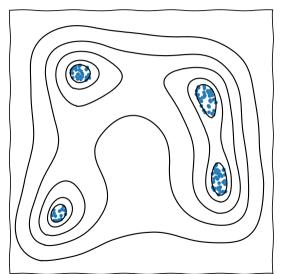


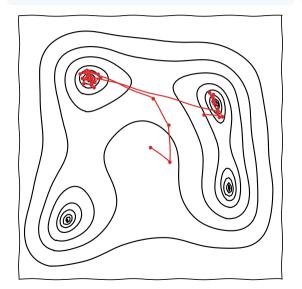


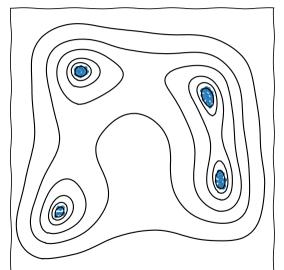


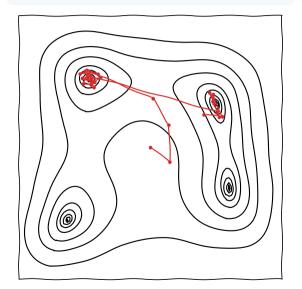


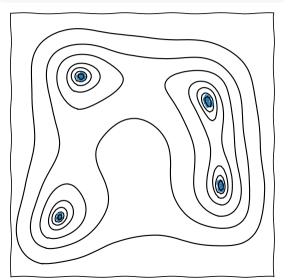


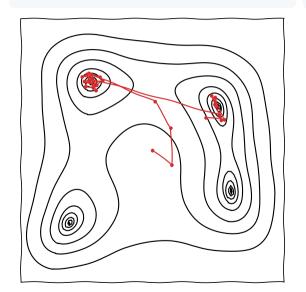


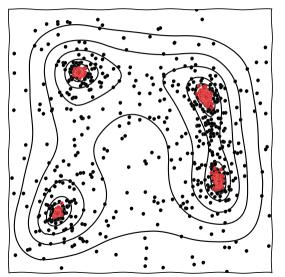




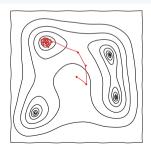




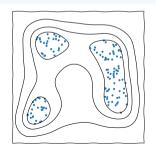




- Single "walker"
- Explores posterior
- ▶ Fast, if proposal matrix is tuned
- Parameter estimation, suspiciousness calculation
- Channel capacity optimised for generating posterior samples

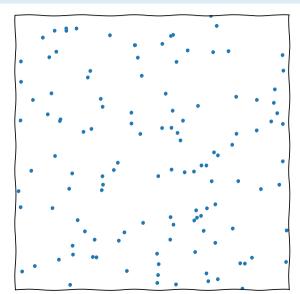


- Ensemble of "live points"
- Scans from prior to peak of likelihood
- Slower, no tuning required
- Parameter estimation, model comparison, tension quantification
- Channel capacity optimised for computing partition function



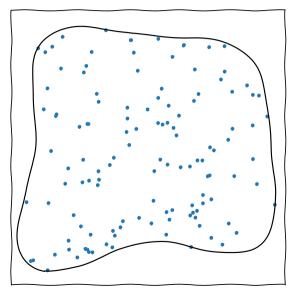
- ▶ Start with *n* random samples over the space.
- Delete outermost sample, and replace with a new random one at higher integrand value.
- ► The "live points" steadily contract around the peak(s) of the function.
- We can use this evolution to estimate volume probabilistically.
- At each iteration, the contours contract by  $\sim \frac{1}{n}$  of their volume.
- ▶ This is an exponential contraction, so

$$\int f(x)dV \approx \sum_{i} f(x_{i})\Delta V_{i}, \quad V_{i} = V_{0}e^{-i/n}$$



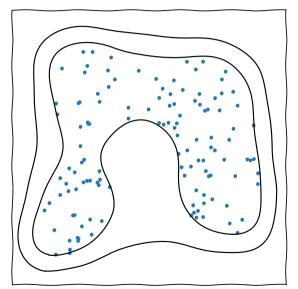
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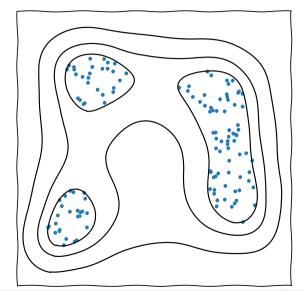
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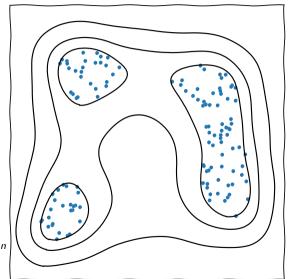
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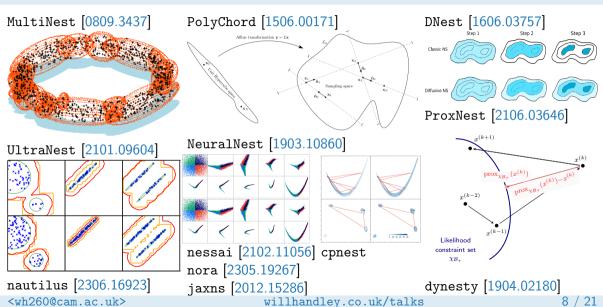


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- At each iteration, the contours contract by  $\sim \frac{1}{n} \pm \frac{1}{n}$  of their volume.
- ▶ This is an exponential contraction, so

$$\int f(x)dV \approx \sum_{i} f(x_i) \Delta V_i, \quad V_i = V_0 e^{-(i \pm \sqrt{i})/n}$$



# Implementations of Nested Sampling [2205.15570] (NatReview)



# Types of nested sampler

- ▶ Broadly, most nested samplers can be split into how they create new live points.
- i.e. how they sample from the hard likelihood constraint  $\{\theta \sim \pi : \mathcal{L}(\theta) > \mathcal{L}_*\}$ .

### **Rejection samplers**

- ▶ e.g. MultiNest, UltraNest.
- ▶ Constructs bounding region and draws many invalid points until  $\mathcal{L}(\theta) > \mathcal{L}_*$ .
- Efficient in low dimensions, exponentially inefficient  $\sim \mathcal{O}(e^{d/d_0})$  in high  $d > d_0 \sim 10$ .

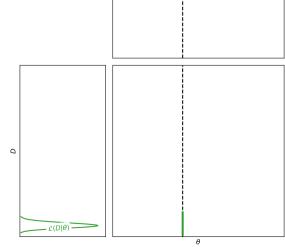
### **Chain-based samplers**

- e.g. PolyChord, ProxNest.
- Run Markov chain starting at a live point, generating many valid (correlated) points.
- Linear  $\sim \mathcal{O}(d)$  penalty in decorrelating new live point from the original seed point.

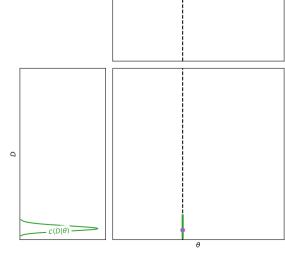
- Nested samplers usually come with:
  - resolution parameter  $n_{\text{live}}$  (which improve results as  $\sim \mathcal{O}(n_{\text{live}}^{-1/2})$ .
  - ▶ set of *reliability* parameters [2101.04525], which don't improve results if set arbitrarily high, but introduce systematic errors if set too low.
  - e.g. Multinest efficiency eff or PolyChord chain length  $n_{\text{repeats}}$ .

- ▶ What do you do if you don't know  $\mathcal{L}(D|\theta)$ ?
- If you have a simulator/forward model  $\theta \to D$  defines an *implicit* likelihood  $\mathcal{L}$ .
- Simulator generates samples from  $\mathcal{L}(\cdot|\theta)$ .
- With a prior  $\pi(\theta)$  can generate samples from joint distribution  $\mathcal{J}(\theta,D) = \mathcal{L}(D|\theta)\pi(\theta)$  the "probability of everything".
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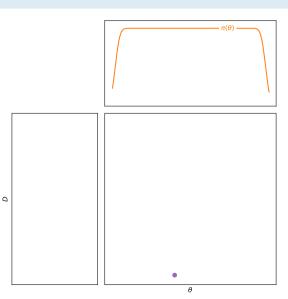
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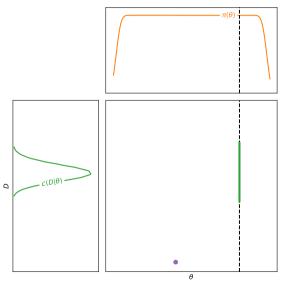
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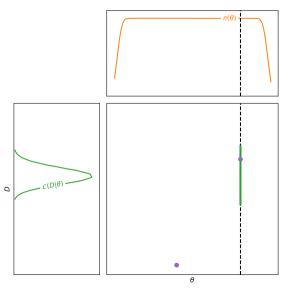
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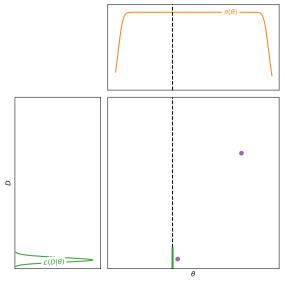
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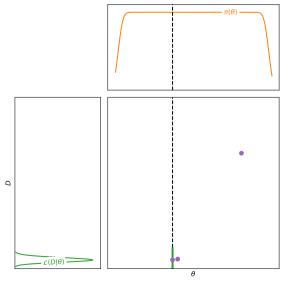
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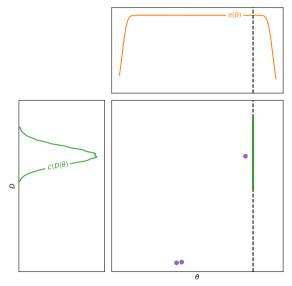
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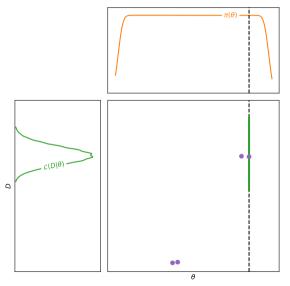
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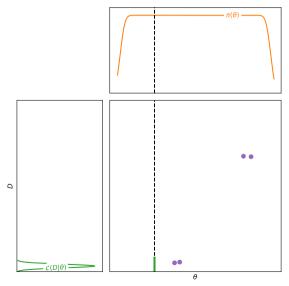
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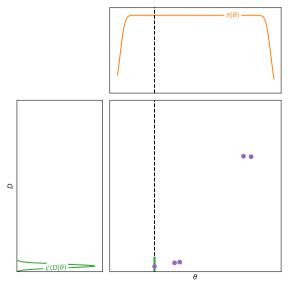
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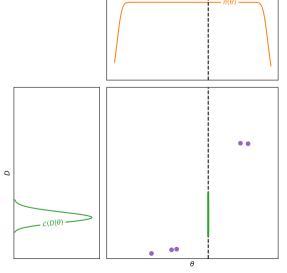
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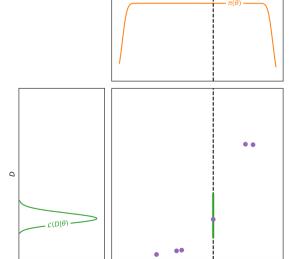
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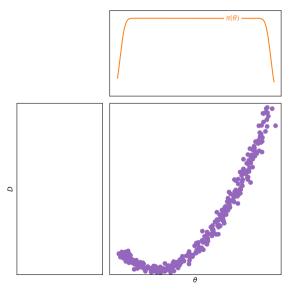
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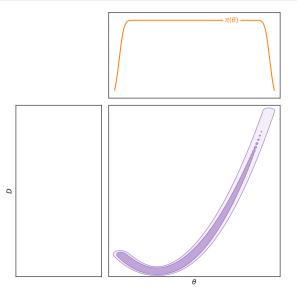
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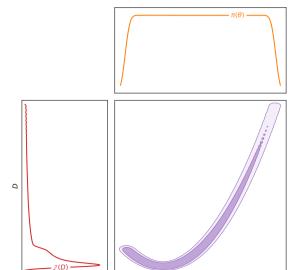
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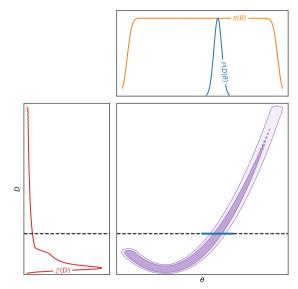
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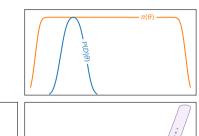
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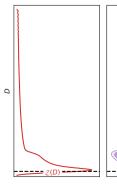


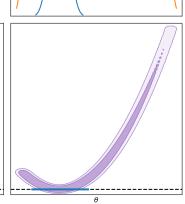
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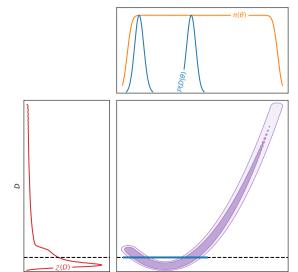
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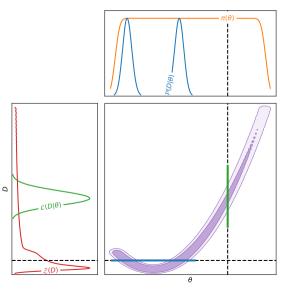




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# Why SBI?

### SBI is useful because:

- If you don't have a likelihood, you can still do inference
  - This is the usual case beyond CMB cosmology
- 2. Faster than LBL
  - emulation also applies to LBI in principle
- 3. No need to pragmatically encode fiducial cosmologies
  - Covariance computation implicitly encoded in simulations
  - ▶ Highly relevant for disentangling tensions & systematics
- 4. Equips AI/ML with Bayesian interpretability
- 5. Lower barrier to entry than LBI
  - Much easier to forward model a systematic
  - Emerging set of plug-and-play packages
  - For this reason alone, it will come to dominate scientific inference



## github.com/sbi-dev



# github.com/undark-lab/swyft



### github.com/florent-leclercq/pyselfi



github.com/justinalsing/pydelfi

# **SBI** in astrophysics

- 2024 has been the year it has started to be applied to real data.
- Mostly for weak lensing
- However: SBI requires mock data generation code
- Most data analysis codes were built before the generative paradigm.
- It's still a lot of work to upgrade cosmological likelihoods to be able to do this (e.g. plik & camspec).

### Investigating the turbulent hot gas in X-COP galaxy clusters

S. Dupourqué<sup>1</sup>, N. Clerc<sup>1</sup>, E. Pointecouteau<sup>1</sup>, D. Eckert<sup>2</sup>, S. Ettori<sup>3</sup>, and F. Vazza<sup>4,5,6</sup>

Dark Energy Survey Year 3 results: simulation-based cosmological inference with wavelet harmonics, scattering transforms, and moments of weak lensing mass maps II. Cosmological property.

M. Gatti, <sup>1, \*</sup> G. Campaille, <sup>2</sup> N. Jeffrey, <sup>3</sup> L. Whiteway, <sup>3</sup> A. Porredon, <sup>4</sup> J. Prat, <sup>5</sup> J. Williamson, <sup>3</sup> M. Raveri, <sup>2</sup> B.

#### Neural Posterior Estimation with guaranteed exact coverage: the ringdown of GW150914

Marco Crisostomi<sup>1,2</sup>, Kallol Dey<sup>3</sup>, Enrico Barausse<sup>1,2</sup>, Roberto Trotta<sup>1,2,4,5</sup>

Applying Simulation-Based Inference to Spectral and Spatial Information from the Galactic Center Gamma-Ray Excess

Katharena Christy,<sup>a</sup> Eric J. Baxter,<sup>b</sup> Jason Kumar<sup>a</sup>

### KiDS-1000 and DES-Y1 combined: Cosmology from peak count statistics

Joachim Hamois-Déraps<sup>1\*</sup>, Sven Heydenreich<sup>2</sup>, Benjamin Giblin<sup>3</sup>, Nicolas Martinet<sup>4</sup>, Tilman Tröster<sup>6</sup>, Marika Asgari<sup>1,6,7</sup>, Pierre Burger<sup>8,0,10</sup>Tiago Castro<sup>1,12,13,14</sup>, Klaus Dolag<sup>15</sup>, Catherine Heymans<sup>3,16</sup>, Hendrik Hildebrandt<sup>16</sup>, Benjamin Joachimi<sup>17</sup> & Angus H. Wright<sup>16</sup>

### KiDS-SBI: Simulation-Based Inference Analysis of KiDS-1000

Maximilian von Wietersheim-Kramsta<sup>1,2,3</sup>, Kiyam Lin<sup>1</sup>, Nicolas Tessore<sup>1</sup>, Benjamin Joachimi<sup>1</sup>, Arthur Loureiro<sup>6,5</sup>, Robert Reischke<sup>6,7</sup>, and Angus H. Wright<sup>7</sup>

# Simulation-based inference of deep fields: galaxy population model and redshift distributions

Beatrice Moser, a.1 Tomasz Kacprzak, a.b Silvan Fischbacher, a Alexandre Refregier, Dominic Grimm, Luca Tortorellic

SMBIG: Cosmological Constraints using Simulation-Based Inference of Galaxy Clustering with Marked Power Spectra

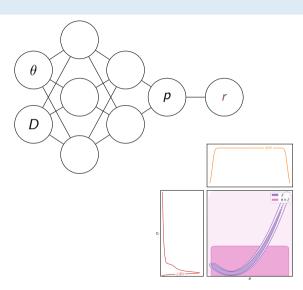
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# **Neural Ratio Estimation**

► SBI flavours: github.com/sbi-dev/sbi

NPE Neural posterior estimation
NLE Neural likelihood estimation
NJE Neural joint estimation
NRE Neural ratio estimation

- ► NRE recap:
  - 1. Generate joint samples  $(\theta, D) \sim \mathcal{J}$ 
    - straightforward if you have a simulator:  $\theta \sim \pi(\cdot)$ ,  $D \sim \mathcal{L}(\cdot|\theta)$
  - 2. Generate separated samples  $\theta \sim \pi$ ,  $D \sim \mathcal{Z}$ 
    - aside: can shortcut step 2 by scrambling the (θ, D) pairings from step 1
  - 3. Train probabilistic classifier p to distinguish whether  $(\theta, D)$  came from  $\mathcal{J}$  or  $\pi \times \mathcal{Z}$ .
  - 4.  $\frac{p}{1-p} = r = \frac{P(\theta, D)}{P(\theta)P(D)} = \frac{\mathcal{J}}{\pi \times \mathcal{Z}} = \frac{\mathcal{L}}{\mathcal{Z}} = \frac{\mathcal{P}}{\pi}$ .
  - 5. Use ratio r for parameter estimation  $\mathcal{P} = r \times \pi$



## **Neural Ratio Estimation**

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    - \* straightforward if you have a simulator:  $\theta \sim \pi(\cdot)$ ,  $D \sim \mathcal{L}(\cdot|\theta)$
  - 2. Generate separated samples  $\theta \sim \pi$ ,  $D \sim Z$ 
    - aside: can shortcut step 2 by scrambling the  $(\theta, D)$  pairings from step 1
  - 3. Train probabilistic classifier p to distinguish whether  $(\theta, D)$  came from  $\mathcal{J}$  or  $\pi \times \mathcal{Z}$ .
  - 4.  $\frac{p}{1-p} = r = \frac{P(\theta, D)}{P(\theta)P(D)} = \frac{\mathcal{J}}{\pi \times \mathcal{Z}} = \frac{\mathcal{L}}{\mathcal{Z}} = \frac{\mathcal{P}}{\pi}$ .
  - 5. Use ratio r for parameter estimation  $\mathcal{P} = r \times \pi$

# **Bayesian proof**

- ▶ Let  $M_{\mathcal{J}}$ :  $(\theta, D) \sim \mathcal{J}$ ,  $M_{\pi \mathcal{Z}}$ :  $(\theta, D) \sim \pi \times \mathcal{Z}$
- ► Classifier gives  $p(\theta, D) = P(M_{\mathcal{J}}|\theta, D) = 1 P(M_{\pi Z}|\theta, D)$
- Bayes theorem then shows  $\frac{p}{1-p} = \frac{P(M_{\mathcal{J}}|\theta,D)}{P(M_{\pi\mathcal{Z}}|\theta,D)} = \frac{P(\theta,D|M_{\mathcal{J}})P(M_{\mathcal{J}})}{P(\theta,D|M_{\pi\mathcal{Z}})P(M_{\pi\mathcal{Z}})} = \frac{\mathcal{J}}{\pi\mathcal{Z}},$ 
  - $P(M_{\mathcal{J}}) = P(M_{\pi \mathcal{Z}}),$
  - and by definition

where we have assumed

 $\bullet \pi(\theta) \mathcal{Z}(D) = P(\theta, D|M_{\pi \mathcal{Z}}).$ 

# **Neural Ratio Estimation**

► SBI flavours: github.com/sbi-dev/sbi

NPE Neural posterior estimation
NLE Neural likelihood estimation
NJE Neural joint estimation
NRE Neural ratio estimation

- NRE recap:
  - 1. Generate joint samples  $(\theta, D) \sim \mathcal{J}$ 
    - straightforward if you have a simulator:  $\theta \sim \pi(\cdot)$ ,  $D \sim \mathcal{L}(\cdot|\theta)$
  - 2. Generate separated samples  $\theta \sim \pi$ ,  $D \sim Z$  aside: can shortcut step 2 by scrambling the
    - $(\theta, D)$  pairings from step 1
  - 3. Train probabilistic classifier p to distinguish whether  $(\theta, D)$  came from  $\mathcal{J}$  or  $\pi \times \mathcal{Z}$ .
  - 4.  $\frac{p}{1-p} = r = \frac{P(\theta, D)}{P(\theta)P(D)} = \frac{\mathcal{J}}{T \times \mathcal{J}} = \frac{\mathcal{L}}{\mathcal{J}} = \frac{\mathcal{P}}{T}$ .
  - 5. Use ratio r for parameter estimation  $\mathcal{P} = r \times \pi$

### Why I like NRE

- The link between classification and inference is profound.
- Density estimation is hard Dimensionless r divides out the hard-to-calculate parts.

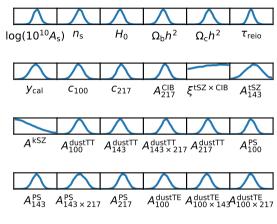
# Why I don't like NRE

- Practical implementations require marginalisation [2107.01214], or autoregression [2308.08597].
- Model comparison and parameter estimation are separate [2305.11241].

# Marginal inference



- Many cosmological likelihoods come with nuisance parameters that have limited relevance for onward inference.
- ▶ Notation: CMB cosmology
  - Likelihood (e.g. plik),
    D Data (e.g. CMB).
  - $\theta$  Cosmological parameters (e.g.  $\Omega_m$ ,  $H_0$ ...),
  - $\alpha$  Nuisance parameters (e.g.  $\Omega_m$ ,  $N_0$ ...),
  - $\alpha$  Nuisance parameters (e.g.  $A_{\text{planck}}$ ...), M Model (e.g.  $\Lambda$ CDM).
- ► Some marginal statistics (e.g. marginal means, posteriors...) are easy to compute.
- More machinery is needed for e.g. nuisance marginalised likelihoods and marginal KL divergences  $\mathcal{D}_{\text{KI}}$ .





# Marginal inference





- Many cosmological likelihoods come with nuisance parameters that have limited relevance for onward inference.
- Notation: GW cosmology

Likelihood

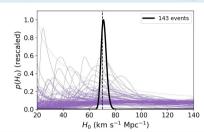
(e.g. LAL), (e.g. GW170817).

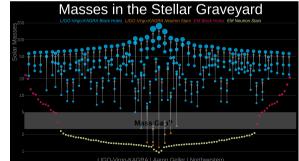
D Data

e.g. GVV170817, (e.g.,  $H_0...$ ),

- $\theta$  Cosmological parameters
- (e.g.,  $m_0$ ...), (e.g.  $m_1$ ,  $m_2$ ...),
- $\alpha$  Nuisance parameters M Model

- (e.g.  $\Lambda$ CDM).
- Some marginal statistics (e.g. marginal means, posteriors...) are easy to compute.
- More machinery is needed for e.g. nuisance marginalised likelihoods and marginal KL divergences D<sub>KL</sub>.





# Nuisance marginalised likelihoods: Theory [2207.11457]



Bayes theorem

$$\mathcal{L}(\theta, \alpha) \times \pi(\theta, \alpha) = \mathcal{P}(\theta, \alpha) \times \mathcal{Z}$$
 (1 Likelihood × Prior = Posterior × Evidence

 $\alpha$ : nuisance parameters,  $\theta$ : cosmo parameters.

Marginal Bayes theorem

$$\mathcal{L}(\theta) \times \pi(\theta) = \mathcal{P}(\theta) \times \mathcal{Z} \tag{2}$$

Non-trivially gives nuisance-free likelihood

$$\mathcal{L}(\theta) = \frac{\mathcal{P}(\theta)\mathcal{Z}}{\pi(\theta)} = \frac{\int \mathcal{L}(\theta, \alpha)\pi(\theta, \alpha)d\alpha}{\int \pi(\theta, \alpha)d\alpha}$$
(3)

### **Key properties**

- Given datasets A and B, each with own nuisance parameters  $\alpha_A$  and  $\alpha_B$ :
- (1) If you use  $\mathcal{L}_A(\theta)$ , you get the same (marginal) posterior and evidence if you had run with nuisance parameters  $\alpha_A$  (ditto B).
  - If you run inference on  $\mathcal{L}_A(\theta) \times \mathcal{L}_B(\theta)$ , you get the same (marginal) posterior and evidence if you had run with all nuisance parameters  $\alpha_A$ ,  $\alpha_B$  on.

(weak marginal consistency requirements on joint  $\pi(\theta, \alpha_A, \alpha_B)$  and marginal priors)

- ▶ To compute the nuisance marginalised likelihood, need: 1. Bayesian evidence Z

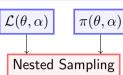
- $\mathcal{L}(\theta, \alpha)$ 
  - $\pi(\theta, \alpha)$

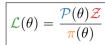
- 2. Marginal prior and posterior densities
- Bayesian evidence  $\mathcal{Z}$ : g
  - Nested sampling
  - Parallel tempering (pocomc, ptmcmc)
  - Sequential Monte Carlo (SMC)
  - MCFvidence
- 2. Marginal prior  $\pi(\theta)$  and posterior  $\mathcal{P}(\theta)$  densities:
  - Histograms of samples
  - Kernel density estimation
  - Normalising flows / Diffusion models
  - •
  - Emulators usually much faster than original likelihoods
  - margarine: PyPI, github.com/htjb/margarine



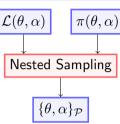
$$\mathcal{L}(\theta) = \frac{\mathcal{P}(\theta)\mathcal{Z}}{\pi(\theta)}$$

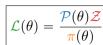
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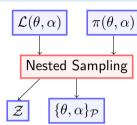


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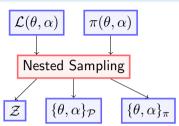




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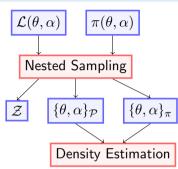


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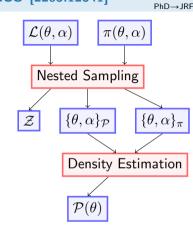


PhD→ IRE

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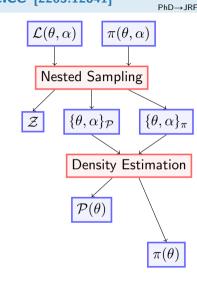
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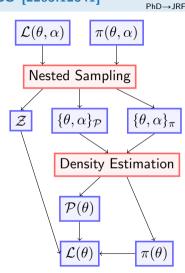
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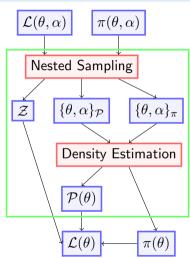
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# Nuisance marginalised likelihoods: Example uses



- Library of pre-trained bijectors to be used as priors/emulators/nuisance marginalised likelihoods (DiRAC allocation unimpeded)
- e.g. easy to apply a *Planck*/DES/HERA/JWST prior or likelihood to your existing MCMC chains without needing to install the whole cosmology machinery.
- Hierarchical modelling:
  - ▶ Usually, have N objects, each with nuisance parameters  $\alpha_i$ , and shared parameters of interest  $\theta$ .
  - Likelihood  $\mathcal{L}(\{D_i\}|\theta,\{\alpha_i\}) = \prod_{i=1}^{N} \mathcal{L}_i(D_i|\theta,\alpha_i)$  has  $N \times \text{len}(\alpha_i) + \text{len}(\theta)$  parameters
  - Instead, break problem down into N runs on  $len(\theta) + len(\alpha_i)$  parameters, and one final one on  $len(\theta)$  parameters, using nuisance marginal likelihoods  $\mathcal{L}_i(\theta)$ .
  - In addition to computational tractability, also can perform model comparison with nuisance marginalised likelihoods.

# The scaling frontier of nested sampling



## How fast in nested sampling?

$$T = T_{\mathcal{L}} imes n_{\mathsf{live}} imes \mathcal{D}_{\mathsf{KL}} imes f_{\mathsf{sampler}}$$

# How accurate is nested sampling?

$$\sigma(\log Z) pprox \sqrt{\mathcal{D}_{\mathsf{KL}}/\mathit{n}_{\mathsf{live}}}$$

in d dimensional parameter space:

$$\mathcal{L}$$
: likelihood eval time  $\sim \mathcal{O}(a)$ 

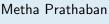
$$n_{\text{live}}$$
: number of live points  $\sim \mathcal{O}(d)$ 

$$\mathcal{D}_{\mathsf{KL}}$$
: KL divergence from prior to posterior  $\approx \log V_{\pi}/V_{\mathcal{P}} \sim \mathcal{O}(d)$ 

$$f_{\text{sampler}}$$
: efficiency of point generation region  $\sim \mathcal{O}(e^{d/d_0})$  or path  $\sim \mathcal{O}(d)$ 

- $T_{\mathcal{L}}$ : likelihood eval time  $\sim \mathcal{O}(d)$  Algorithmically improving  $f_{\text{sampler}}$  is only a fraction of the story!
  - $\triangleright \mathcal{D}_{KI}$  appears twice, so improvements here are quadratically important.
  - Gradients give you d more information.

 $ightharpoonup T \sim \mathcal{O}(d^4)$  whilst polynomial is far from ideal, athough progress can be made on all fronts.





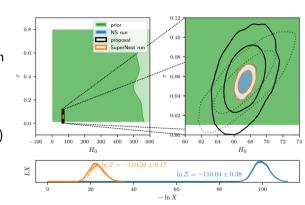
For constant "run quality"  $\sigma$ ,

$$T = T_{\mathcal{L}} \times n_{\text{live}} \times \mathcal{D}_{\text{KL}} \times f_{\text{sampler}}, \quad \sigma \approx \sqrt{\mathcal{D}_{\text{KL}} / n_{\text{live}}}$$

$$\Rightarrow \boxed{T = T_{\mathcal{L}} \times \sigma \times \mathcal{D}_{\text{KL}}^2 \times f_{\text{sampler}}}$$

so if you can reduce the KL divergence, then quadratic gains to be made

- This can be crudely achieved by choosing a narrower prior  $\pi^*$  and then correcting the evidence  $\mathcal{Z} = \mathcal{Z}^* \frac{V_\pi^*}{V_\pi}$  (REACH [2210.07409])
- This can be made more sophisticated with SuperNest [2212.01760] & posterior repartitioning
- Recent application to gravitational waves by Metha Prathaban Ongoing work



# Jax-based nested samplers



- very recent work over the past month
- ► Have implemented a nested slice sampler in blackjax [#755]
- 1 pip install git+https://github.com/handley-lab/blackjax@nested\_sampling 2 import blackjax.ns.adaptive
- parallelised over vmapped likelihood & prior evaluations
- ▶ Plugs into jim [kazewong/jim] and ripple [2302.05329]
- interested in finding use-cases for such a sampler this week
- Also interested in understanding current limitations/strengths of jax/GPU GW programming

### Conclusions



github.com/handley-lab

- ▶ **Next-generation inference:** Addressing challenges in astrophysical analysis, whether likelihoods are available or not.
- ▶ **Simulation-based inference (SBI):** Leveraging simulations and machine learning for posterior estimation and model comparison when only a forward model exists. Advantages include speed, implicit covariance handling, and ease of incorporating systematics.
- Marginal statistics: Efficient computation of nuisance-marginalised likelihoods using margarine. Applications to hierarchical modeling and building prior/likelihood libraries.
- ► Accelerated nested sampling: Improving the scaling of nested sampling with techniques like beta flows, posterior repartitioning, and SuperNest.
- ▶ Jax-based nested samplers: Introducing recent work on parallel nested sampling implementations in Jax for increased performance.

# **SGW** relevant portions

- Marginal inference
  - ▶ Earth term/pulsar term from Stas Babak is ripe for a margarine based approach.
  - Linked to the SBI approach for GWB global fit that Bryan Zaldivar is working on.
- Nested sampling is an (underused) alternative for computing Bayes factors in the PTA community
  - current approaches are often Savage-Dicke density ratio/PTMCMC parallel tempering.
- Global fitting could benefit from accelerated approaches
  - jax-based GW codes are currently in development.
  - ▶ ML offers many opportunities for acceleration (emulation, proposals, . . . )